



## Occupation and lymphoid malignancies: results from a French case-control study.

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## Abstract

**Objectives:** Investigating relationships between potential occupational risk factors and lymphoid malignancy (LM).

**Methods:** We conducted a multicenter hospital-based case-control study in France between 2000-2004, including 824 incident cases of non-Hodgkin's lymphoma (NHL), Hodgkin's lymphoma (HL), multiple myeloma (MM) and 'lymphoproliferative syndrome' (LPS) and 752 frequency-matched controls. Data were collected through face-to-face standardized and detailed interviews.

**Results:** *Farming* was significantly associated with NHL (OR=1.4 [1.0-2.0]) and, although not significantly, with LPS and MM. ORs were higher for longest durations of exposure. Self-declared exposure to *pesticides* was significantly associated with NHL (OR=1.8 [1.2-2.7]) and HL (OR=2.2 [1.0-4.7]). Neither solvent-related jobs nor self-reported exposure to solvents were related to LM. Systematic screening based on job titles did not evidence any other association.

**Conclusions:** The results support the hypothesis that farming plays a role in most types of LM.

## INTRODUCTION

The expression 'lymphoid malignancies' (LM) covers various diseases affecting lymphoid tissue, namely: Hodgkin's lymphoma (HL), non-Hodgkin's lymphoma (NHL), multiple myeloma (MM) and lymphoproliferative syndrome (LPS). In France, LM are the most frequent cancers after smoking-related cancer and the number of incident LM cases was estimated to be approximately 17,000 in 2000. The annual number of deaths was almost 9,000.(1) In recent decades, the incidence of NHL has dramatically increased in almost all industrialized countries(2, 3) and data from the French registries show a yearly rate of increase that was greater than 3% (3.8% in men, 3.5% in women) between 1978 and 2000.(1) The causes of the increase in NHL remain largely unexplained. The increase began before the AIDS epidemic and does not seem to be fully explained by changes in diagnostic methods or registration.(3, 4) Changes in lifestyle and occupation are thus expected to play a role in the increase.

The etiology of LM remains largely unknown, except for a few established risk factors. Thus, congenital and acquired immunodeficiencies are well documented predisposing factors for NHL;(5) infection by Epstein Barr virus (EBV) is strongly related to Burkitt's lymphoma and HL;(6) and *Helicobacter pylori* infections are associated with MALT lymphoma.(7) Malaria has been identified as a cofactor of EBV-induced Burkitt's lymphoma in Africa, but the other cofactors, particularly in western countries, have yet to be discovered. In addition to the known risk factors, occupational exposures, particularly in agriculture-related jobs, are strongly suspected of being risk factors for LM.

Numerous studies have investigated for an association between farming and NHL. The case-control studies conducted in the USA,(8-13) Italy,(14-16) France(17) and Spain,(18) have almost all reported some degree of positive association with farming. A meta-analysis estimated that the excess NHL risk associated with farming was weak, but slightly more marked in the USA.(19)

MM has also been shown to be associated with farming-related jobs(14, 20-24) and possibly with glyphosate application.(25) For LPS, the association has been less well documented. Nonetheless, a French study(26) and a Swedish study(27) have reported associations between hairy cell leukemia, a rare LPS entity, and farming. Farmers are subject to a variety of exposures, including solvents, various organic antigens, microbial infections and pesticides. Several pesticides have been found to be associated with increased LM risk, including carbamates,(28-30) and organophosphates,(31, 32) as have phenoxyacetic acid herbicides,(29, 30, 33, 34) DDT(30, 34, 35) and organochlorine insecticides, aldrin,(30) dieldrin,(12) chlordane,(9) lindane(30, 36, 37) and toxaphene.(12)

This study investigated the role of lifelong occupations in the occurrence of the main ICD-O categories of LM using data from a large-scale, multicenter, case-control study.

## MATERIALS AND METHODS

### Study population

A French multicenter hospital-based case-control study was conducted in the main hospitals of Bordeaux, Brest, Caen, Nantes, Lille and Toulouse from September 2000 to December 2004. The eligible cases were incident cases with a diagnosis of lymphoid malignancy (LM), aged between 18 and 75 years, and residing in the hospital catchment area of each center. The diagnosis of LM was documented by cytology and histology, and reviewed by a team of pathologists. All cases were classified using the WHO classification (ICD-O-3). Cases treated with immunosuppressant drugs and cases with AIDS were not eligible. Except for the LPS cases, who were included up to 18 months post-diagnosis in view of the good prognosis, all cases had to be recruited within 6 months of diagnosis. Most cases (88.9%) were included within 3 months. Of the 872 subjects eligible during the recruitment period, 48 (5.5%) refused to participate. The study population thus consisted in 824 incident cases of LM, classified using ICD-O-3, and further subdivided into four categories: HL (n = 149, ICD-O-3 codes (9650-9655/3, 9659/3, 9661-9665/3, 9667/3)), NHL (n = 399 consisting of 173 cases of diffuse large B cell-lymphoma (DLCL) (9679/3, 9680/3), 101 cases of follicular lymphoma (FL) (9690/3, 9691/3, 9695/3, 9698/3), 21 cases of lymphoplasmacytic lymphoma/Waldenstrom macroglobulinemia (9671/3, 9761/3), 17 cases of marginal zone B-cell lymphoma of the MALT type (9699/3), 3 cases of splenic marginal zone B-cell lymphoma (9689/3), 25 cases of T-cell lymphoma (9702/3, 9705/3, 9714/3, 9729/3), 25 cases of mantle-cell lymphoma (9673/3), and 34 cases of other lymphoma (9728/3, 9687/3, 9826/3, 9591/3)), MM (n = 108, (9731-9732/3)) and LPS (n = 168, 132 cases of chronic lymphocytic leukemia (CLL) (9823/3, 9670/3) and 36 cases of hairy cell leukemia (HCL) (9940/3)).

The hospital-based design of the study was chosen since case and control blood samples were required. Special care was therefore paid to selecting an appropriate control group. The controls were patients with no prior history of hematological malignancy recruited in the same hospitals as the cases, mainly in orthopedic and rheumatological departments. Subjects admitted for cancer or a disease directly related to occupation (occupational injuries), smoking (chronic obstructive pulmonary disease, ischemic cardiovascular disease) or alcohol consumption (alcoholic cirrhosis) were not eligible as controls in order to avoid over-representation of some of the factors of interest. The controls were individually matched with the cases by center, age ( $\pm 3$  years) and gender. The aim of matching was to ensure that at least one control would be available for each case. Out of the 853 eligible controls identified during the recruitment period, 100 refused the interview (11.7%) and 1 subject whose interview was incomplete was excluded *a posteriori*. The final control sample thus consisted of 752 subjects. The reasons for hospital admission of the controls were mainly orthopedic or rheumatological (fractures (19.9%), wounds (0.9%), other non-occupational injuries (10.8%), osteoarthritis (23.3%), back diseases (15.3%), polyarticular diseases (4.0%), infectious bone and joint diseases (2.9%), minor musculoskeletal malformations (4.8%), other diseases of the bones and joints (6.2%)), peripheral nervous disorders (2.1%), digestive,

urinary or genital tract diseases (4.8%), cardiovascular diseases (1.2%), skin and subcutaneous tissues diseases (1.9%), and infections (0.8%).

## Data collection

The study protocol was submitted to the CNIL (90003) and DGS (2000/0107) for data privacy and ethical approval. Data collection was conducted in two stages. The case and control patients first completed a standardized self-administered questionnaire eliciting information on their socioeconomic characteristics, familial medical history, and lifetime residential and occupational histories. Regarding occupational history, each job held for at least six months was to be reported, with the job title, workplace name and business, job start and end dates, and a description of the specific tasks and products personally handled (open-ended question). The patients were also asked to report occupational exposures to a broad predefined list of agents (paints, adhesives, petrol, wood preservatives, pesticides, ionizing radiation, etc.) they could have been exposed to, and to state their average exposure frequency ('never', 'occasionally', 'at least once/week', 'every day').

The patients then underwent a face-to-face interview (average duration: 80 minutes) by trained staff using a structured standardized questionnaire eliciting personal and familial medical histories, lifestyle characteristics (smoking and alcohol, tea and coffee consumption), outdoor leisure activities and non-occupational exposures. At the end of the interview, the self-administered questionnaire was reviewed with the interviewer and, if necessary, further information was elicited using *ad hoc* occupational questionnaires specific to potential exposures, which will undergo further expert review by industrial hygienists.

Blood samples were obtained from the cases and controls after consent form signature and the biological specimens (sera, constitutional DNA, tumour tissue) were placed on storage.

## Variables analyzed

Job titles and industries were encoded using the 1968 edition of the International Standard Classification of Occupations issued by the International Labour Organization (ILO) and the first revision of the Statistical Classification of Economic Activities in the European Community (NACE), respectively, by a trained coder, blind with respect to case/control status. Socioeconomic categories at the time of inclusion were generated from the last job held and encoded using the ILO code at the 2-digit level (0.1 to 2.1: 'scientific and administrative managers', 3.0 to 5.2: 'administrative, sales and service workers', 5.3 to 5.9 or 7.1 to 9.9: 'factory workers', 6.0 to 6.4: 'agricultural workers'), and were also divided into white collar (ILO codes 0.1 to 5.2) and blue collar (ILO codes 5.3 to 9.9) categories. Subjects having worked at least 6 month in a specific job or industry were considered exposed to that job or industry, and those who had never worked in that specific subgroup constituted the baseline category. The total duration of employment in a specific job or industry was obtained by summing the durations of all job periods in that particular job or industry. A categorical variable (unexposed, duration <10

years, duration  $\geq 10$  years) was used in the analysis. Lastly, all the subjects who reported exposure, at any level, to one of the agents in the broad predefined list, were considered exposed to that agent.

## Data analysis

The pair-matching used as a basis for the recruitment was broken in order to enable the whole control group to be used for the analysis of all LM types, with stratification by age (5-year age groups), gender and center. For each subgroup, HL, NHL, LPS and MM, the control group consisted in all the controls who could be included in one of the strata covered by the corresponding subgroup of cases.

The analysis of occupations consisted in 3 stages: (1) systematic screening for the occupations related to LM, which tested all the categories of occupation (2-digit ILO codes for jobs and 2-digit NACE codes for activities), with at least ten exposed subjects and at least one exposed case and one exposed control; (2) a specific analysis of farming and jobs involving solvent exposure, both corresponding to specific *a priori* hypotheses based on the literature. The duration of employment (never, duration  $<10$  years, duration  $\geq 10$  years) was used to quantify the exposures and time lags were applied to account for latency; (3) analysis of the occupational exposures derived from the self-declared checklist.

All the analyses were conducted using SAS software version 9.1 (SAS institute, Cary NC., 1989). Odds ratios (OR) and their 95% confidence intervals (95% CI) were estimated using unconditional logistic regression models including the stratification variables age, gender and center as categorical variables. The Wald test was used to test for linear trends. The duration of employment was incorporated in the model as a quantitative variable after subjects in a given category (0,  $>0$ - $<10$ ,  $\geq 10$  years) had been allocated the median duration of employment in that category.

Analyses were conducted separately for the LM subgroups (Hodgkin's lymphoma, non-Hodgkin's lymphoma, lymphoproliferative syndrome, multiple myeloma), by gender, and for all LM taken together.

In order to check the robustness of the results, conditional logistic regressions restricted to the paired case-control samples were conducted. Polytomous logistic regressions with a 5-level nominal non-ordered response variable (Control, HL cases, NHL cases, LPS cases, MM cases) in which the common comparator was the entire control group, were also conducted.

## Study power

For NHL, with power of 80% and a two-sided alpha error of 5%, the size of the study sample was sufficient to evidence ORs between 1.5 and 2.7 for exposures with prevalences ranging from 2% to 20%. For other subtypes of LM (HL, LPS, MM), ORs between 2.0 and 4.0 could be evidenced for the corresponding exposure prevalence.

## RESULTS

The distribution of the cases and controls by stratification variable is shown in table 1. The use of the whole control group assigned more than 2 controls to each case in most strata, except in the youngest categories, in which HL predominated. In the latter, there was therefore a significant age difference between the HL cases and the controls. The MM cases differed from the controls with regard to gender since they did not show the male predominance observed in other lymphoid malignancies. Lastly, significant differences were also observed for the centers, mainly because the Caen hospital had a higher LPS recruitment than the other centers.

With regard to socio-demographic characteristics (Table 2), the cases and controls were similarly distributed with respect to socioeconomic category, urban/rural residential status, educational level, number of jobs held and duration of employment, except for the HL cases, who were less often factory workers and had a higher educational level than the controls.

### Screening

Systematic analysis of the occupations showed significantly elevated ORs for 'farmers' (ILO code 6.1) associated with all groups of LM except LPS (OR [95%CI]: 1.8 [1.0-3.0], 2.9 [1.0-8.0] and 2.6 [1.3-5.5] for NHL, HL and MM, respectively). 'Agricultural and animal husbandry workers' (ILO 6.2) were associated with LPS (OR = 1.7 [1.0-2.8]) and NHL (OR = 1.5 [1.1-2.2]). 'Transport equipment operators' (ILO 9.8) were negatively associated with NHL (OR = 0.6 [0.4-1.0]) and positively associated with LPS (OR = 1.9 [1.1-3.4]), while employment as 'cooks, waiters, bartenders' (ILO 5.3) was only associated with MM (OR = 2.2 [1.0-4.7]). Screening based on industrial subgroups revealed no association for HL. The 'manufacture of chemicals and chemical products' sector (NACE 2 digit code 24) was positively linked with NHL (OR = 2.1 [1.1-4.0]) in contrast to the 'land transport; transport via pipelines' sector (NACE 60), which was negatively linked to NHL (OR = 0.6 [0.3-1.0]). A negative association between 'activities of households as employers of domestic staff' (NACE 95) and LPS was noted (OR = 0.3 [0.1-0.8]). The following industrial sectors showed significant positive associations with MM: 'agriculture, hunting and related service activities' (NACE 01); 'manufacture of wearing apparel; dressing and dyeing of fur' (NACE 18); 'insurance and pension funding' (NACE 66) (ORs: 1.8 [1.0-3.1], 2.5 [1.0-6.0] and 8.4 [2.2-32.8], respectively). Conversely, the 'other business activities' sector (NACE 74) was negatively linked with MM (OR = 0.2 [0.1-0.9]).



## Specific hypotheses

### Agriculture-related jobs

The cases had been employed in agriculture-related jobs more often than the controls (table 3). The associations tended to be more marked for employment durations greater than 10 years and for farmers (ILO 6.1), rather than for agricultural workers. For NHL, the association was significant (OR = 1.4 [1.0-2.0]) and was observed for both genders, with ORs increasing with the duration of exposure (*p for trend* = 0.02). For men, significant associations were also observed with both MM and HL when the exposure duration was greater than 10 years. Similar patterns were observed for agricultural activities (NACE 01) (Table 3). The associations remained when different hypothetical lag times were considered, i.e., when the last 10, 20, 30 or 40 years of employment before diagnosis or interview were considered unexposed. The associations with LPS even seemed to be strengthened as the latency period increased.

### Occupations involving potential solvent exposure

The associations between solvent-related jobs and LM are shown separately for men and women in table 4. For men, the only significant association was between NHL and employment in NACE 24: 'manufacture of chemicals and chemical products' (OR = 2.6 [1.2-5.9]). For women, the occupations: 'building caretakers, char workers or cleaners' (ILO 5.51 and 5.52) were negatively associated with all LM, and significantly negatively associated with NHL. When all the above occupations were grouped together as solvent-exposed jobs, no association was evidenced, even for long term ( $\geq 10$  years) exposure. The numbers were too small to investigate long latency periods for most occupations, but, when the numbers were sufficient, applying various lag times did not enable any association to be evidenced.

## Self-declared exposure

Table 5 reports the associations between LM and exposures from the checklist included in the self-administered questionnaire. Exposure to pesticides used for crops was significantly associated with all LM but LPS, with higher ORs for more frequent exposures. The use of wood preservatives was also associated with NHL. The associations between LM and farming were very close with (82 cases, 61 controls; OR = 1.6 [1.1-2.3]) and without (84 cases, 64 controls; OR = 1.5 [1.0-2.2]) self-declared exposure to pesticides, compared to subjects who had declared no exposure to pesticides and no agriculture related job. This applied to both men and women. In contrast, self-declared exposure to pesticides with no agriculture-related job was markedly associated with LM (25 cases, 11 controls; OR = 2.5 [1.2-5.1]).

Exposure to 'paints, lacquers, varnishes' was negatively associated with LPS. No trend was observed with increasing frequency of exposure (OR: 0.5 [0.3-0.9] for occasional exposure and 0.5 [0.2-1.1] for exposure at least once per week). In order to investigate the impact of missing data, the analysis was repeated with the missing data considered no exposure for the cases and frequent exposure for the controls. The results were unchanged. The opposite classification

of the missing data (missing data considered frequent exposure for the cases and no exposure for the controls) failed to reveal any association.

## Adjustments and sensitivity analysis

Adjustment for urban/rural residential status, educational level, duration of employment, and number of jobs held did not change any of the results. Consistent results were obtained using conditional models, in which the LM subgroups had independent control groups, and using polytomous models. This accounts for the use of a common control group for all LM groups. Finally, the effects on the findings of each center and each reason for control admission were investigated. The results remained stable when either the centers or the reason-for-admission categories were excluded from the analysis one by one.

## DISCUSSION

The main result of the present study consists in the positive associations found between most LM groups and employment in agriculture-related jobs for at least six months. The association was significant for NHL (OR = 1.4 [1.0-2.0]) and of the same order of magnitude, although on the borderline of significance, for LPS and MM. For HL cases, the association was limited to farmers (OR = 2.9 [1.0-8.0]). Generally, associations were strengthened for exposures of duration greater than 10 years, with a significant trend for NHL. The cases reported self-declared exposure to pesticides more often than the controls and the associations were more marked for the highest reported exposure frequencies. No association with occupational exposure to solvents was observed.

Preferential selection of farmer cases was limited since the cases were recruited in main hospitals unlikely to attract specific patient categories. Moreover, only cases living in the hospital catchment areas were recruited in order to limit potential selection on educational and socioeconomic level. The inclusion of cases was systematic and blind to the patients' occupations and educational levels. The refusal rate was low (5.5%). In addition, there is no obvious reason for selection being similar in all the centers and the results were shown to be rugged in sensitivity analyses. Lastly, over-representation of farmers among the cases through selection of survivors is unlikely since inclusion took place shortly (median: 41 days) after diagnosis.

The controls were mainly recruited in orthopedic and rheumatological departments. Admissions for cancer or diseases directly related to smoking (chronic obstructive pulmonary disease, ischemic cardiovascular disease), alcohol intake (alcoholic cirrhosis) or occupation were excluded, in order to avoid artificial over-representation of some risk factors or socioeconomic categories among the controls. However, there was no restriction of control selection on past medical history, which might have under-represented particular jobs or lifestyles in the controls. Moreover, the distribution of

socioeconomic status in the control group was very similar to that of the overall French population as determined by the National Institute for Statistics and Economic Studies in 2003.(38)

Residual selection of controls was nonetheless possible and, for that reason, we verified that the associations remained stable when groups of controls with particular reasons for admission were excluded from the analysis one-by-one. Residual selection could also have been dependent on the center but the associations were also stable after exclusion of each center from the analysis.

Lastly, although it is very credible that all LM may share common risk factors, the similar associations with agricultural professions observed for most LM subtypes may have been due to the common control group. However, the results were very similar when conditional analyses, in which each control was assigned to only one case, were conducted.

The cases and controls were interviewed in the hospitals, under the same conditions, by the same trained interviewers for each center and using standardized questionnaires, thus reducing the scope for differential misclassifications. In addition, neither the subjects nor the interviewers were aware of the specific hypotheses tested, the study being presented to both as broadly related to 'environment and health', and the person responsible for job history coding was blind to case-control status.

Non-differential misclassifications due to defective recall, common in retrospective studies, probably do not play a major role in the present results based on job history, which is more objective and easier to recall than tasks or contacts with products. The details given in the questionnaire helped code jobs and industries, and the person who coded occupations had extensive experience with the codes and coding process. However, job titles may be poor surrogates for the exposures of interest, particularly since they were aggregated at the 2-digit level. An association between solvent exposure and LM cannot be ruled out on the basis of the present results, particularly if exposures are uncommon or if associations are weak.

The analyses were performed with the stratification variables (age, gender and center) included in the models, and separately, by gender. Additional adjustments for urban/rural residence status, educational level, duration of employment and number of jobs held did not modify the results.

When different lag-time scenarios were applied, the relationship between farming-related jobs and LM was found not to be restricted to the recent exposure periods. However, most of the recent exposures had started early, and the variance of period of employment, conditionally on age and duration of employment, was too small to enable reliable conclusions.

The screening analysis evidenced very few associations other than with agriculture-related activities, which were specifically under study, and there was no particular consistency between NACE and ILO categories. The associations may have been observed by chance, since multiple tests were run with a view to generating hypotheses.

Self-reported information on occupational exposure is probably not as reliable as information drawn from specialized interviews administered by trained interviewers. Exposures may be underreported, although frequent or long-lasting exposures are probably easier to recall than occasional exposures. Broad categories of exposure, as elicited by the self-administered questionnaire, may lead to excessively sensitive subject classification. Non-differential misclassifications are therefore probable and may have contributed to the absence of association observed with self-reported exposures to solvents. However, the result is consistent with the absence of association observed with jobs involving solvent exposures. Non-differential misclassifications may also explain the absence of any association with pesticide exposure self-reported by farmers, although an association was observed for non-farmers. Farmers may have overlooked occasional exposures that were unremarkable in a context of common pesticide use by others.

The statistical power of the study was limited for some types of LM. The minimum detectable odds ratios for risk factors with a prevalence of about 30%, like that of solvent-related jobs, were 1.7, 1.4, 1.8, 1.9, for LPS, NHL, HL and MM, respectively. Nevertheless, the estimates were close to unity and no dose-response pattern was evidenced, making the lack of power unlikely to explain the absence of association with solvent-related jobs or self-declared exposures to solvents.

The relationship between LM and occupational exposures has been investigated in many countries, with various definitions of the diseases and occupations. Overall, the results based on mortality or morbidity suggested associations, but were largely unable to elucidate the role of pesticides, antigen stimulation and animal viruses in those associations. The IARC decision to classify non-arsenical insecticide application, considered as a whole, as probably carcinogenic (group 2A) reflects the complexity of singling out a particular agent employed in farming activities.<sup>(39)</sup> Meta-analyses have estimated ORs greater than unity for agricultural occupations and NHL,<sup>(19)</sup> MM<sup>(40)</sup> and HL.<sup>(41)</sup> More recent publications on studies conducted in the USA,<sup>(13)</sup> Italy,<sup>(16)</sup> Germany<sup>(21)</sup> and Spain<sup>(18)</sup> have reported heterogeneous results with regard to the association between lifelong occupational history and LM. The pooled analysis<sup>(13)</sup> of two previous case-control studies conducted in Kansas and Nebraska<sup>(8, 10)</sup> reported significant associations, for men, between NHL and CLL taken as a whole and employment as a farmer (OR = 1.6 [1.2-2.1]) or an agricultural activity (OR = 1.9 [1.4-2.6]). Analyses by LM subtype indicated that the increase in risk was particularly marked for CLL (OR = 5.7 [3.1-10.7] for farmers and 8.9 [4.5-17.3] for agricultural activity). In the German study, employment as a 'farmer' (ILO 6.0-6.1) was also associated with LM, but the relationship was finally restricted to MM (OR = 9.2 [2.6-33.1]), based on a limited number of exposed subjects.<sup>(21)</sup> The Spanish data showed no association with 'ever farming', but when the farmers

where subdivided into the types of farming jobs performed (e.g. crop farming, animal farming, general farming), significantly increased ORs were observed for 'general farming' for all LM subtypes. In contrast, Costantini *et al.* reported no association with farming.(16) The present results showed ORs of about 1.5 for all LM combined and for specific LM subtypes. The ORs are of the same order of magnitude as the estimates generated by meta-analyses. The higher estimates for farmers than for agricultural workers may be real and due to the fact that, in France where mixed farming predominates, pesticide use was usually considered an expert task that farm owners preferred not to delegate, as has been previously observed in our previous study on hairy cell leukemia.(26) A stronger association for farmers was also reported by Mester *et al.*.(21) The present results suggest a positive trend with the duration of employment in farming. This has also been reported by Zheng *et al.* ,(13) while Mester *et al.*(21) observed the highest ORs for farming durations of less than 10 years. With respect to the exposure to pesticides itself, the current analysis is only based on the self-declared use of pesticides on crops, wood preservatives and treatment of buildings. Interestingly, the associations were more marked for the highest exposure frequencies.

## Conclusion

The results of this study, based on occupational history and self-reported exposures, support the hypothesis that farming plays a role in most types of LM but do not evidence any association with solvent-related occupations. They also suggest that pesticides may be involved in the association.

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## Competing interests

None.

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**Table 1:** Comparability of cases and controls according to stratification variables (LPS: lymphoproliferative syndrome; NHL: non-Hodgkin lymphoma; HL: Hodgkin's lymphoma; MM: multiple myeloma; LM: lymphoid malignancies)

	LPS		NHL		HL		MM		Controls		Cases
	Cases n=168 %	Controls n=464 %	Cases n=399 %	Controls n=701 %	Cases n=149 %	Controls n=417 %	Cases n=108 %	Controls n=478 %	Cases n=752 %		
Gender	ns										
Female	38	34	39	38	42	36	48	35	39	40	ns
Male	62	66	61	62	58	64	52	65	61	60	
Age (y)	ns										
18-24	0	0	3	4	23	9	0	0	5	5	ns
25-29	0	0	2	3	13	8	0	0	4	4	
30-34	0	0	4	4	11	7	1	0	4	4	
35-39	3	4	6	6	11	8	2	2	6	5	
40-44	5	4	8	8	12	9	3	3	8	7	
45-49	8	10	10	9	9	13	8	9	9	9	
50-54	13	14	17	17	7	18	18	24	16	15	
55-59	21	18	17	13	3	9	14	15	13	15	
60-64	17	20	12	15	3	8	22	19	14	13	
65-69	20	16	11	11	7	10	20	15	11	13	
70 +	14	14	11	10	1	1	12	12	10	10	
Center	***										
Bordeaux	10	16	26	23	21	21	24	24	21	22	ns
Brest	13	24	22	20	15	18	19	22	19	18	
Caen	38	19	6	12	11	12	9	11	13	14	
Lille	18	14	8	9	3	2	17	11	10	10	
Nantes	15	15	19	17	27	25	11	12	18	19	
Toulouse	5	12	19	20	23	22	19	19	19	17	

ns: non significant; \*:p<0.05; \*\*:p<0.01; \*\*\*:p<0.001

**Table 2:** Comparability of cases and controls by socio-demographic and employment characteristics. ORs were estimated by unconditional logistic regression models including the stratification variables: age, gender and center (LPS: lymphoproliferative syndrome; NHL: non-Hodgkin's lymphoma; HL: Hodgkin's lymphoma; MM: multiple myeloma; LM: lymphoid malignancies)

	LPS						NHL						HL						MM						all LM					
	Ca			Co			Ca			Co			Ca			Co			Ca			Co			Ca			Co		
	OR	[95%CI]	ns	OR	[95%CI]	ns	OR	[95%CI]	ns	OR	[95%CI]	*	OR	[95%CI]	ns	OR	[95%CI]	ns	OR	[95%CI]	ns	OR	[95%CI]	ns	OR	[95%CI]	ns	OR	[95%CI]	ns
Socioeconomic categorie	46	111	1.0	ref.			98	169	1.0	ref.			44	101	1.0	ref.			26	108	1.0	ref.			214	185	1.0	ref.		
Administrative, sales / service workers	39	96	1.1	[0.7-1.9]			90	142	1.1	[0.7-1.6]			38	78	1.4	[0.8-2.5]			24	100	1.2	[0.6-2.2]			191	149	1.1	[0.8-1.5]		
Scientific / administrative managers	67	210	0.8	[0.5-1.2]			158	317	0.9	[0.6-1.2]			38	186	0.5	[0.3-0.9]			41	218	0.9	[0.5-1.5]			304	335	0.8	[0.6-1.0]		
Factory workers	15	39	1.2	[0.6-2.4]			42	47	1.6	[1.0-2.6]			8	26	1.0	[0.4-2.6]			14	38	2.2	[1.0-4.8]			79	48	1.4	[1.0-2.2]		
Agricultural workers	1	8	0.4	[0.1-3.8]			11	26	0.8	[0.3-1.8]			21	26	1.0	[0.4-2.1]			3	14	0.7	[0.2-2.8]			36	35	0.8	[0.5-1.5]		
Unemployed																														
White collar/blue collar <sup>1</sup>																														
White collar	85	207	1.0	ref.			188	311	1.0	ref.			82	179	1.0	ref.			50	208	1.0	ref.			405	334	1.0	ref.		
Blue collar	82	249	0.8	[0.5-1.1]			200	364	0.9	[0.7-1.2]			46	212	0.5	[0.3-0.8]			55	256	1.0	[0.6-1.5]			383	383	0.8	[0.7-1.0]		
Unemployed	1	8	0.4	[0.1-3.6]			11	26	0.7	[0.3-1.6]			21	26	0.8	[0.4-1.8]			3	14	0.7	[0.2-2.5]			36	35	0.8	[0.5-1.4]		
Urban/rural status of the place of residence																														
> 5000 inhabitants	113	293	1.0	ref.			251	447	1.0	ref.			101	258	1.0	ref.			70	297	1.0	ref.			535	487	1.0	ref.		
≤ 5000 inhabitants	55	171	0.9	[0.6-1.3]			148	254	1.0	[0.8-1.4]			48	159	0.9	[0.6-1.3]			38	181	1.0	[0.7-1.7]			289	265	1.0	[0.8-1.2]		
Education																														
None or primary level	69	178	1.0	ref.			134	221	1.0	ref.			17	101	1.0	ref.			39	175	1.0	ref.			259	232	1.0	ref.		
Lower secondary level	57	170	0.9	[0.6-1.4]			136	263	0.8	[0.6-1.2]			48	159	1.3	[0.7-2.5]			37	174	1.1	[0.7-1.9]			278	281	0.9	[0.7-1.2]		
Upper secondary or tertiary level	42	116	1.2	[0.7-1.9]			129	217	1.0	[0.7-1.4]			84	157	1.8	[1.0-3.5]			32	129	1.3	[0.7-2.2]			287	239	1.2	[0.9-1.5]		
Number of jobs held																														
0	1	8	0.5	[0.1-4.1]			11	26	0.8	[0.4-1.7]			21	26	1.1	[0.5-2.4]			3	14	0.7	[0.2-2.6]			36	35	0.9	[0.5-1.5]		
[1-2]	41	119	1.0	ref.			127	221	1.0	ref.			56	138	1.0	ref.			33	133	1.0	ref.			257	235	1.0	ref.		
[3-4]	52	163	1.0	[0.6-1.6]			133	249	0.9	[0.7-1.3]			44	146	0.9	[0.6-1.5]			37	177	1.0	[0.6-1.7]			266	262	0.9	[0.7-1.2]		
5 +	74	174	1.1	[0.6-1.8]			128	205	1.2	[0.8-1.7]			28	107	0.9	[0.5-1.6]			35	154	1.2	[0.7-2.2]			265	220	1.1	[0.9-1.5]		
Duration of employment (y)																														
<24	24	85	1.0	ref.			109	236	1.0	ref.			107	194	1.0	ref.			20	87	1.0	ref.			260	268	1.0	ref.		
[24-37]	60	178	1.4	[0.7-2.7]			160	251	1.6	[1.0-2.3]			29	151	0.8	[0.4-1.6]			42	198	1.3	[0.6-2.5]			291	259	1.3	[1.0-1.9]		
>37	84	201	2.1	[1.0-4.2]			130	214	1.6	[1.0-2.5]			13	72	1.2	[0.4-3.2]			46	193	1.5	[0.7-3.2]			273	225	1.5	[1.0-2.2]		

ns: non-significant; \*: p<0.05; \*\*: p<0.01; <sup>1</sup>White collar: ILO 0.1 to 5.2; Blue collar: ILO 5.3 to 9.9

**Table 3:** Association between LM and farming. ORs are estimated by unconditional logistic regression models including stratification variables age, gender, center and socioeconomic category (LPS: lymphoproliferative syndrome; NHL: non-Hodgkin lymphoma; HL: Hodgkin's lymphoma; MM: multiple myeloma; LM: lymphoid malignancies)

Occupations (ILO)	LPS				NHL				HL				MM				all LM			
	Ca/Co	OR	[95% CI]	p	Ca/Co	OR	[95% CI]	p	Ca/Co	OR	[95% CI]	p	Ca/Co	OR	[95% CI]	p	Ca/Co	OR	[95% CI]	p
Men																				
Agricultural, Animal Husbandry / Forestry Workers (6)																				
< 10 years	33/77	1.4	[0.8-2.4]		59/92	1.5	[1.0-2.3]		15/56	1.5	[0.7-3.2]		19/71	2.2	[1.1-4.6]	*	126/94	1.6	[1.1-2.2]	**
≥ 10 years	17/34	1.6	[0.8-3.2]		28/45	1.4	[0.8-2.4]		6/34	0.8	[0.3-2.2]		8/28	2.3	[0.9-6.0]		59/47	1.5	[1.0-2.3]	
Farmers (ILO 6.1)	16/43	1.2	[0.6-2.5]		31/47	1.6	[0.9-2.8]		9/22	3.3	[1.2-8.9]	*	11/43	2.2	[0.9-5.2]		67/47	1.7	[1.1-2.6]	*
≥ 10 years	9/23	1.2	[0.5-2.8]		19/25	1.6	[0.8-3.1]		7/13	4.4	[1.5-13.5]	**	10/23	3.6	[1.5-8.7]	**	45/25	2.0	[1.2-3.4]	**
Agricultural / animal husbandry workers (6.2)	30/62	1.5	[0.8-2.7]		51/73	1.7	[1.1-2.7]	*	11/41	1.4	[0.6-3.2]		14/55	1.7	[0.8-3.6]		106/75	1.6	[1.1-2.4]	**
Women																				
Agricultural, Animal Husbandry / Forestry Workers (6)																				
< 10 years	10/22	1.8	[0.7-4.8]		25/28	1.5	[0.8-2.7]		2/15	0.5	[0.1-2.7]		7/21	1.0	[0.4-2.7]		44/32	1.4	[0.8-2.3]	
≥ 10 years	4/12	1.3	[0.4-4.6]		8/17	0.8	[0.3-2.0]		2/10	0.7	[0.1-3.6]		3/12	0.8	[0.2-3.0]		17/19	0.9	[0.5-1.9]	
Farmers (ILO 6.1)	6/10	2.8	[0.7-11]		17/11	2.5	[1.1-5.8]	*	0/5	-	-		4/9	1.3	[0.3-4.9]		27/13	2.1	[1.0-4.4]	*
≥ 10 years	4/7	1.9	[0.4-8.6]		12/8	2.3	[0.9-5.9]		0/3	-	-		3/7	1.2	[0.3-5.4]		19/10	1.8	[0.8-4.2]	
Agricultural / animal husbandry workers (6.2)	8/18	1.9	[0.7-5.5]		19/22	1.4	[0.7-2.9]		2/12	0.7	[0.1-3.4]		5/17	0.8	[0.3-2.6]		34/26	1.3	[0.7-2.3]	
Both																				
Agricultural, Animal Husbandry / Forestry Workers (6)																				
< 10 years	43/99	1.6	[1.0-2.6]		84/120	1.4	[1.0-2.0]	*	17/71	1.1	[0.6-2.2]		26/92	1.6	[0.9-2.9]		170/126	1.5	[1.1-2.0]	**
≥ 10 years	21/46	1.6	[0.9-2.9]		36/62	1.2	[0.7-1.8]		8/44	0.7	[0.3-1.7]		11/40	1.5	[0.7-3.2]		76/66	1.3	[0.9-1.8]	
Farmers (ILO 6.1)	22/53	1.6	[0.8-2.9]		48/58	1.7	[1.1-2.7]	*	9/27	2.2	[0.9-5.4]		15/52	1.8	[0.9-3.6]		94/60	1.8	[1.2-2.6]	**
≥ 10 years	13/30	1.4	[0.7-3.0]		31/33	1.8	[1.0-3.0]	*	7/16	2.9	[1.0-8.0]	*	13/30	2.6	[1.3-5.5]	*	64/35	2.0	[1.3-3.0]	**
Agricultural / animal husbandry workers (6.2)	38/80	1.7	[1.0-2.8]	*	70/95	1.5	[1.1-2.2]	*	13/53	1.1	[0.6-2.2]		19/72	1.4	[0.7-2.5]		140/101	1.5	[1.1-2.1]	**
Activities (NACE)																				
Men																				
Agriculture, hunting / related service activities (01)																				
< 10 years	30/71	1.3	[0.7-2.2]		53/81	1.5	[1.0-2.3]		11/46	1.4	[0.6-3.0]		19/64	2.3	[1.1-4.7]	*	113/83	1.6	[1.1-2.2]	*
≥ 10 years	16/35	1.3	[0.6-2.8]		28/43	1.4	[0.8-2.4]		4/32	0.6	[0.2-1.9]		9/29	2.3	[0.9-5.8]		57/45	1.4	[0.9-2.3]	
Farmers (ILO 6.1)	14/36	1.2	[0.6-2.5]		25/38	1.6	[0.9-2.9]		7/14	3.9	[1.3-11.6]	*	10/35	2.3	[1.0-5.6]		56/38	1.7	[1.1-2.7]	*
Women																				
Agriculture, hunting / related service activities (01)																				
< 10 years	12/22	2.2	[0.9-5.6]		26/30	1.4	[0.8-2.6]		3/16	0.6	[0.1-2.3]		8/21	1.2	[0.5-3.2]		49/33	1.5	[0.9-2.5]	
≥ 10 years	5/10	2.1	[0.6-7.1]		9/16	0.9	[0.4-2.3]		3/8	1.0	[0.2-4.4]		3/11	0.9	[0.2-3.5]		20/17	1.2	[0.6-2.4]	
Farmers (ILO 6.1)	7/12	2.3	[0.7-7.7]		17/14	1.9	[0.9-4.3]		0/8				5/10	1.6	[0.5-5.5]		29/16	1.8	[0.9-3.6]	
Both																				
Agriculture, hunting / related service activities (01)																				
< 10 years	42/93	1.6	[1.0-2.5]		79/111	1.4	[1.0-2.0]		14/62	1.0	[0.5-1.9]		27/85	1.8	[1.0-3.1]	*	162/116	1.5	[1.1-2.0]	**
≥ 10 years	21/45	1.6	[0.9-2.9]		37/59	1.2	[0.8-1.9]		7/40	0.7	[0.3-1.6]		12/40	1.6	[0.8-3.4]		77/62	1.3	[0.9-1.9]	
Farmers (ILO 6.1)	21/48	1.6	[0.8-2.9]		42/52	1.6	[1.0-2.6]	*	7/22	1.7	[0.7-4.4]		15/45	2.0	[1.0-4.0]		85/54	1.7	[1.2-2.5]	**

\*, p<0.05, \*\*, p<0.01

**Table 4:** Relationship between lymphoid malignancies (LM) and jobs considered to involve exposure to organic solvents for men and women. ORs were estimated by unconditional logistic regression models including the stratification variables: age, center and socioeconomic category (white collar/blue collar) (LPS: lymphoproliferative syndrome; NHL: non-Hodgkin's lymphoma; HL: Hodgkin's lymphoma; MM: multiple myeloma)

Job titles (ILO codes)	LPS				NHL				HL				MM				all LM			
	ca/co	OR	[95% CI]	p	ca/co	OR	[95% CI]	p	ca/co	OR	[95% CI]	p	ca/co	OR	[95% CI]	p	ca/co	OR	[95% CI]	p
Men																				
Painters (8.19.55; 8.73.70; 9.3)	5/15	0.7	[0.2-2.2]		14/22	1.3	[0.6-2.6]		7/17	2.1	[0.8-5.8]		3/15	1.1	[0.3-4.2]		29/23	1.4	[0.8-2.4]	
Artists and designers (1.61-1.62)	0/1	-	-		3/1	4.5	[0.4-45.9]		1/0	-	-		0/1	-	-		4/1	3.1	[0.3-28.8]	
Printers (9.2)	3/3	2.1	[0.4-11.8]		2/6	0.6	[0.1-3.1]		0/1	-	-		1/5	0.6	[0.0-7.0]		6/7	0.8	[0.2-2.3]	
Mechanics (8.43-8.44-8.49)	11/50	0.6	[0.3-1.2]		25/57	0.8	[0.5-1.3]		3/34	0.4	[0.1-1.3]		5/46	0.5	[0.2-1.4]		44/61	0.7	[0.4-1.0]	
Electrical fitters (8.5)	10/17	2.0	[0.8-4.7]		22/29	1.5	[0.8-2.8]		2/16	0.4	[0.1-1.9]		6/19	2.2	[0.8-5.9]		40/29	1.4	[0.9-2.4]	
Machine tool operators (8.32-8.41)	10/22	1.1	[0.5-2.5]		13/24	1.1	[0.6-2.4]		2/15	0.6	[0.1-3.1]		4/20	1.0	[0.3-3.4]		29/26	1.1	[0.6-1.9]	
Cabinet makers, joiners (8.1; 9.54)	7/19	0.8	[0.3-2.2]		13/24	1.1	[0.5-2.2]		5/16	1.4	[0.4-4.5]		3/19	1.0	[0.3-3.8]		28/25	1.1	[0.6-2.0]	
Rubber and plastic product makers (9.0)	3/2	2.4	[0.4-15.7]		1/2	1.3	[0.1-16.0]		0/1	-	-		1/1	6.4	[0.3-116.0]		5/2	2.4	[0.5-12.6]	
Shoe and leather goods makers (8.0)	0/3	-	-		4/4	2.2	[0.5-9.4]		0/0	-	-		0/3	-	-		4/4	1.2	[0.3-4.8]	
Chemical industry (NACE 24)	2/9	0.6	[0.1-3.1]		15/12	2.6	[1.2-5.9]	*	1/8	0.4	[0.0-4.1]		2/7	1.0	[0.2-6.3]		20/13	1.6	[0.8-3.2]	
Any of the above solvent-related jobs	39/121	0.8	[0.5-1.3]		95/159	1.3	[0.9-1.8]		18/93	0.7	[0.4-1.4]		22/116	1.1	[0.6-2.1]		174/165	1.1	[0.8-1.4]	
< 10 years	14/54	0.6	[0.3-1.1]		42/75	1.2	[0.8-2.0]		10/47	0.7	[0.3-1.7]		8/53	0.9	[0.4-2.1]		74/78	1.0	[0.7-1.4]	
≥ 10 years	25/67	0.9	[0.5-1.7]		53/84	1.3	[0.9-2.1]		8/46	0.7	[0.3-1.8]		14/63	1.3	[0.6-2.7]		100/87	1.2	[0.8-1.7]	
Women																				
Housekeepers (5.5)	9/36	0.7	[0.2-1.9]		17/52	0.3	[0.2-0.7]	**	5/29	0.5	[0.1-1.6]		10/36	0.7	[0.3-1.8]		41/59	0.5	[0.3-0.8]	**
Nurses (0.71-0.72)	3/7	0.8	[0.2-3.9]		8/9	1.8	[0.6-5.0]		1/2	1.2	[0.1-17.9]		3/9	0.9	[0.2-4.1]		15/10	1.2	[0.5-2.7]	
Chemical industry (NACE 24)	1/5	0.3	[0.0-2.9]		6/8	1.5	[0.5-4.5]		1/6	0.5	[0.0-4.5]		0/4	-	-		8/8	0.9	[0.3-2.6]	
Any solvent-related job	16/51	0.7	[0.3-1.4]		39/73	0.8	[0.5-1.3]		9/38	0.7	[0.2-1.8]		12/51	0.6	[0.3-1.2]		76/81	0.7	[0.5-1.1]	
< 10 years	7/26	0.4	[0.2-1.3]		14/37	0.6	[0.3-1.2]		4/22	0.4	[0.1-1.6]		5/21	0.6	[0.2-2.0]		30/42	0.6	[0.3-1.0]	*
≥ 10 years	9/25	0.9	[0.4-2.3]		25/36	1.0	[0.5-1.8]		5/16	1.2	[0.3-4.4]		7/30	0.5	[0.2-1.4]		46/39	0.9	[0.5-1.5]	

\*, p<0.05; \*\*, p<0.01

**Table 5:** Relationship between lymphoid malignancies (LM) and self-declared occupational exposures. ORs were estimated by unconditional logistic regression models including the stratification variables: age, gender, center and socioeconomic category (white collar/blue collar) (LPS: lymphoproliferative syndrome; NHL: non-Hodgkin's lymphoma; HL: Hodgkin's lymphoma; MM: multiple myeloma)

	LPS				NHL				HL				MM				all LM			
	ca/co	OR	[95% CI]	p	ca/co	OR	[95% CI]	p	ca/co	OR	[95% CI]	p	ca/co	OR	[95% CI]	p	ca/co	OR	[95% CI]	p
Ionizing radiation	7/25	0.8	[0.3-2.0]		11/27	0.7	[0.3-1.5]		5/14	1.3	[0.4-3.9]		2/19	0.6	[0.1-2.5]		25/29	0.8	[0.4-1.3]	
Dyeing for wood, textiles, skins	11/41	0.7	[0.3-1.4]		29/54	1.0	[0.6-1.7]		11/25	1.4	[0.6-3.2]		3/36	0.4	[0.1-1.4]		54/55	0.9	[0.6-1.4]	
Paints, lacquers, varnishes	28/113	0.5	[0.3-0.8]	**	78/132	1.2	[0.8-1.7]		27/74	1.5	[0.8-2.7]		16/96	0.8	[0.4-1.6]		149/139	1.0	[0.8-1.4]	
Glues, adhesives, cements	35/105	0.8	[0.5-1.4]		74/142	1.0	[0.7-1.4]		29/77	1.4	[0.8-2.5]		17/93	0.9	[0.5-1.8]		155/144	1.1	[0.8-1.4]	
Thinners or pickling solutions	39/123	0.8	[0.5-1.2]		86/161	1.0	[0.7-1.4]		25/89	0.9	[0.5-1.6]		21/113	0.9	[0.5-1.7]		171/167	1.0	[0.7-1.3]	
Grease-removers and cleaners for metals, glass, etc.	44/122	0.9	[0.6-1.5]		84/162	1.0	[0.7-1.3]		30/89	1.3	[0.7-2.2]		18/104	0.9	[0.5-1.8]		176/166	1.0	[0.8-1.3]	
Solvents used as chemicals reagents	11/38	0.7	[0.3-1.5]		27/46	1.1	[0.7-1.9]		9/22	1.5	[0.6-3.6]		4/28	0.7	[0.2-2.2]		51/46	1.1	[0.7-1.6]	
Petrol	32/95	0.9	[0.5-1.4]		62/120	0.9	[0.6-1.4]		18/60	1.0	[0.5-2.0]		17/84	1.2	[0.6-2.4]		129/124	1.0	[0.7-1.4]	
Wood preservatives	14/42	1.0	[0.5-1.9]		40/55	1.5	[1.0-2.4]		10/28	1.3	[0.6-3.1]		8/37	1.3	[0.5-3.0]		72/55	1.3	[0.9-2.0]	
Pesticides for crops	20/59	1.1	[0.6-2.1]		60/69	1.8	[1.2-2.7]	**	14/36	2.2	[1.0-4.5]	*	13/54	1.3	[0.6-2.6]		107/72	1.6	[1.1-2.2]	**
occasionally	17/50	1.1	[0.6-2.1]		46/60	1.6	[1.0-2.4]	*	11/31	2.0	[0.9-4.4]		8/47	0.9	[0.4-2.1]		82/62	1.4	[1.0-2.0]	
at least once/week	3/9	1.3	[0.3-5.3]		14/9	3.6	[1.5-8.6]	*	3/5	3.3	[0.7-15.6]		5/7	3.9	[1.1-13.4]	*	25/10	2.7	[1.3-5.7]	*
Treatment of buildings against insects	7/30	0.6	[0.3-1.5]		18/34	1.0	[0.5-1.8]		6/21	1.3	[0.5-3.5]		3/26	0.6	[0.2-2.1]		34/36	0.9	[0.6-1.5]	

\*: p<0.05, \*\*: p<0.01